

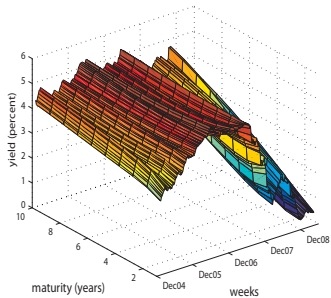
# On Correlation and Default Clustering in Credit Markets

Antje Berndt\*, Peter Ritchken and Zhiqiang Sun

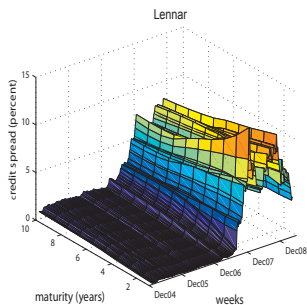
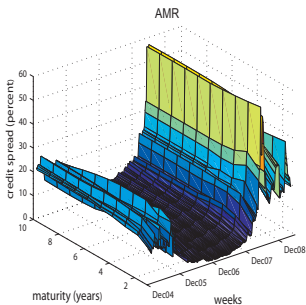
\*Tepper School of Business  
Carnegie Mellon University

Bachelier Congress, Toronto 2010

Panel A: Riskless Yield Curves



Panel B: Credit Spread Curves



# Single-Name Credit Risk Pricing – What do we do?

Develop general yet tractable Markovian HJM models that

- fully incorporate information on riskless and credit spread term structures

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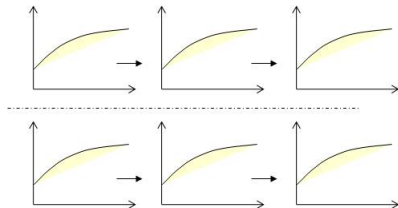
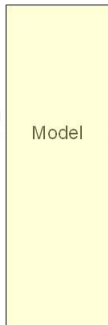
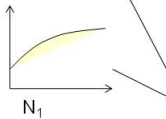
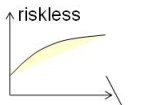
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- allow different volatility structures for forward rates, that can be initialized to closely match empirical structures
- credit spreads and yield curves are represented by a finite set of state variables
- allow arbitrary interest rate-credit spread correlations
- permit shocks to the economy to impact riskless yield curves and credit spreads



# Multi-Name Credit Risk Pricing – What do we do?

Extend single-name Markovian HJM models to

- multi-name infection-type models



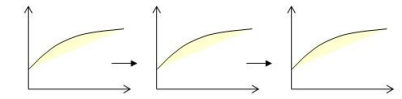
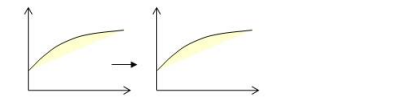
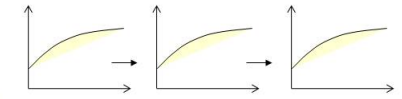
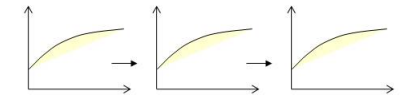
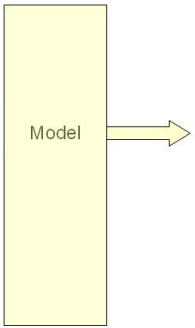
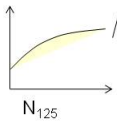
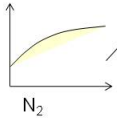
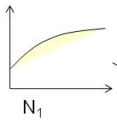
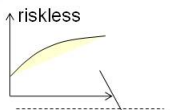
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Using Kalman filter parameter estimates, we show the importance of

- interest rate-credit spread correlations
- default contagion
- the initial credit spread curve distribution



## HJM Models: Riskless Dynamics

- Let  $P(t, T)$  be the price at date  $t$  of a pure riskless discount bond that pays \$1 at date  $T$ :

$$P(t, T) = e^{-\int_t^T f(t,u)du},$$

where  $f(t, u)$  represents the date- $t$  forward rate for the future time increment  $[u, u + dt]$ .

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- We assume

$$df(t, T) = \mu_f(t, T)dt + \sigma_f(t, T)dz_f(t) + c_f(t, T)dN_f(t),$$

given  $f(0, T)$ .  $N_f(t)$  is independent Poisson process with intensity  $\eta_f$ .

## Dynamics of Riskless Bond Prices

Apply Ito's lemma for jump-diffusion processes to obtain

$$\begin{aligned} \frac{dP(t, T)}{P(t, T)} = & \left( r(t) + \frac{1}{2} \sigma_p(t, T) \sigma'_p(t, T) - \int_t^T \mu_f(t, u) du \right) dt \\ & - \sigma_p(t, T) dz_f(t) + \left( e^{-K_p(t, T)} - 1 \right) dN_f(t), \end{aligned}$$

where

$$\sigma_p(t, T) = \int_t^T \sigma_f(t, u) du,$$

$$K_p(t, T) = \int_t^T c_f(t, u) du.$$

## HJM Models: Risky Debt

- For firm  $A$  that has not defaulted prior to date  $t$ , we have

$$dY_A(t) = \begin{cases} 1 & \text{with probability } \eta_A(X_t)dt \\ 0 & \text{with probability } 1 - \eta_A(X_t)dt, \end{cases}$$

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- The date- $t$  price of a bond issued by  $A$  is given by

$$\Pi_A(t, T) = V_A(t, T)1_{\tau_A > t},$$

where

$$\begin{aligned} V_A(t, T) &= e^{-\int_t^T (f(t, u) + \lambda_A(t, u)) du} \\ &= P(t, T) S_A(t, T). \end{aligned}$$

$\lambda_A(t) = \eta_A(t)\ell_A(t)$  is firm  $A$ 's forward credit spread,  $\eta_A(t)$  is the default arrival intensity, and  $\ell_A(t)$  denotes LGD.

# Credit Spreads Dynamics

We assume

$$d\lambda_A(t, T) = \mu_A(t, T) dt + \sigma_A(t, T) dz_A(t) + c_{fA}(t, T) dN_f(t), t \leq \tau_A,$$

where

- correlation with diffusive riskless term structure:

$$E(dz_f(t) dz'_A(t)) = \Sigma_{m \times n}^A dt = \left( \rho_{ij}^A \right) dt$$

- a jump in riskless rates could transmit to shocks in the credit spreads
- $\sigma_A(t, T)$  is predictable
- $c_{fA}(t, T)$  is a deterministic function of time to maturity,  $T - t$



# Proposition 1: HJM Restrictions on the Drift Terms

No arbitrage implies

$$\begin{aligned}\mu_f(t, T) &= \sigma_p(t, T)\sigma'_f(t, T) - c_f(t, T)e^{-K_p(t, T)}\eta_f \\ \mu_A(t, T) &= \sigma_{S_A}(t, T)\sigma'_A(t, T) + \sigma_f(t, T)\Sigma^A\sigma'_{S_A}(t, T) \\ &\quad + \sigma_p(t, T)\Sigma^A\sigma'_A(t, T) + g_A(t, T),\end{aligned}$$

where

$$g_A(t, T) = \eta_f \left( c_f(t, T)e^{-K_p(t, T)} - (c_f(t, T) + c_{fA}(t, T))e^{-(K_p(t, T) + K_{fA}(t, T))} \right).$$

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Problem with HJM models:

- In general, the dynamics are not Markovian in a small number of state variables
- To overcome this issue, we curtail the volatility structures

# Markovian HJM: Volatilities and Jump-Impact Factors

- **Volatilities** are given by

$$\sigma_{f_i}(t, T) = h_{f_i}(t) e^{-\kappa_{f_i}(T-t)},$$

$$\sigma_{A_j}(t, T) = h_{A_j}(t) e^{-\kappa_{A_j}(T-t)}.$$

where  $h_{f_i}(t)$  and  $h_{A_j}(t)$  are predictable functions.

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- Example:  $h_f(t) = \min(|\tilde{h}_f(t)|, \bar{h}_f)$ , where  $\bar{h}_f$  is a large yet finite constant, and

$$\tilde{h}_f(t) = \sigma_f r(t)$$

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$$\tilde{h}_f(t) = \sigma_f r(t)$$

- **Jump-impact factors** are of the form

$$c_f(t, T) = c_f e^{-\gamma_f(T-t)},$$

$$c_{fA}(t, T) = c_{fA} e^{-\gamma_{fA}(T-t)}$$

## Proposition 2: Markovian Models for Riskless Debt

Under volatility restrictions, we get **exponential affine riskless bond prices**:

$$P(t, T) = \frac{P(0, T)}{P(0, t)} \exp \left( - \sum_{i=1}^2 \sum_{j=1}^m H_{ij}(t, T) \psi_{ij}(t) - H_3(t, T) \psi_3(t) + H_J(t, T) \right),$$

where

$$H_{1j}(t, T) = 1/\kappa_{f_j}^2 \left( 1 - e^{-\kappa_{f_j}(T-t)} \right), \text{ for } j = 1, \dots, m$$

$$H_{2j}(t, T) = -1/(2\kappa_{f_j}^2) \left( 1 - e^{-2\kappa_{f_j}(T-t)} \right), \text{ for } j = 1, \dots, m$$

$$H_3(t, T) = c_f/\gamma_f \left( 1 - e^{-\gamma_f(T-t)} \right),$$

$$H_J(t, T) = \eta_f e^{-\frac{c_f}{\gamma_f}} \int_t^T \left( e^{\frac{c_f}{\gamma_f}} e^{-\gamma_f(u-t)} - e^{\frac{c_f}{\gamma_f}} e^{-\gamma_f u} \right) du.$$

The dynamics of the state variables are

$$d\psi_{1j}(t) = (h_{f_j}^2(t) - \kappa_{f_j} \psi_{1j}(t))dt + \kappa_{f_j} h_{f_j}(t) dz_{f_j}(t)$$

$$d\psi_{2j}(t) = (h_{f_j}^2(t) - 2\kappa_{f_j} \psi_{2j}(t))dt$$

$$d\psi_3(t) = -\gamma_f \psi_3(t)dt + dN_f(t).$$

## Proposition 2: Markovian Models for Risky Debt

Under volatility restrictions, and assuming RMV, the risky bond price at  $t$  is  $\Pi_A(t, T) = P(t, T)S_A(t, T)1_{\tau_A > t}$ , where  $S_A(t, T)$  is exponential affine:

$$\begin{aligned} S_A(t, T) = & \frac{S_A(0, T)}{S_A(0, t)} \exp[-A_0(t, T) - \sum_{j=1}^n (K_{0,j}(t, T)\xi_{0,j} - K_{1,j}(t, T)\xi_{1,j}) \\ & + \sum_{i=1}^m \sum_{j=1}^n (K_{2,ij}(t, T)\xi_{2,ij} - K_{3,ij}(t, T)\xi_{3,ij} - K_{4,ij}(t, T)\xi_{4,ij}) \\ & - K_5(t, T)\xi_5(t)]. \end{aligned}$$

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$$d\xi_{0,j}(t) = (h_{A_j}^2(t) - \kappa_{A_j}\xi_{0j}(t)) dt + \kappa_{A_j}h_{A_j}(t) dz_{A_j}(t)$$

$$d\xi_{1,j}(t) = (h_{A_j}^2(t) - 2\kappa_{A_j}\xi_{1j}(t)) dt$$

$$d\xi_{2,ij}(t) = (h_{f_i}(t)h_{A_j}(t) - (\kappa_{A_j} + \kappa_{f_i})\xi_{2,ij}(t)) dt$$

$$d\xi_{3,ij}(t) = (h_{f_i}(t)h_{A_j}(t) - \kappa_{f_i}\xi_{3,ij}(t)) dt$$

$$d\xi_{4,ij}(t) = (h_{f_i}(t)h_{A_j}(t) - \kappa_{A_j}\xi_{4,ij}(t)) dt$$

$$d\xi_5(t) = -\gamma_{fA}\xi_5(t) + dN_f(t).$$

# Stochastic Drivers vs State Variables

- Assume forward rates are driven by  $m$  **stochastic drivers**, and credit spreads by  $n$ 
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- Assume forward rates are driven by  $m$  **stochastic drivers**, and credit spreads by  $n$ 
  - Computational burden is limited to that of  $(m + n)$  -dim affine models with jumps
- Number of **state variables**:  $3mn + 2(m + n + 1)$
- Number of state variables can sometimes be reduced:
  - $m = n = 1$ : 8
  - $m = n = 1$ , no jumps and constant  $h(\cdot)$  functions: 2
  - $m = n = 1$ , no jumps and no correlations between interest rates and credit spreads: 4

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  - If the time partitions are refined to weeks, the number of state variables increases to **2,880**.
- Markovian HJM:
  - A maximum of **8** state variables need to be maintained, no matter what the partition.

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- This can easily be accommodated in our multifactor models ( $m > 1, n > 1$ ).
- Proposition 2 can be generalized to enable humped volatility structures even when  $m = n = 1$ .
- In fact, we are able to establish arbitrary shapes:

$$\sigma_f(t, T) = h_f(t) \sum_{j=1}^k a_j e^{-\kappa_j(T-t)}, \quad k > 1.$$

## Relationship with Duffie-Kan Affine Models

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## Relationship with Duffie-Kan Affine Models

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- The drift terms of the path statistics offset spot rate volatilities in a manner that allows bond yields to be affine in the states, even though the state variables themselves do not have to be affine processes.

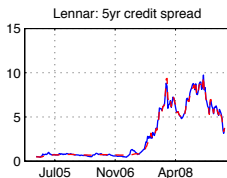
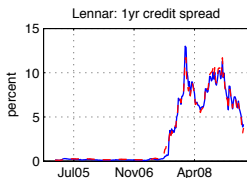
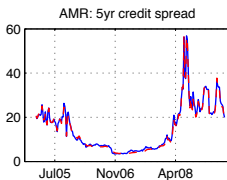
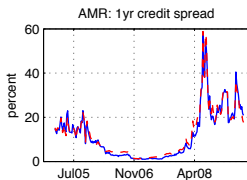
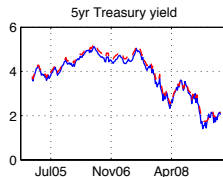
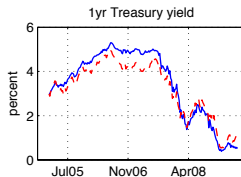
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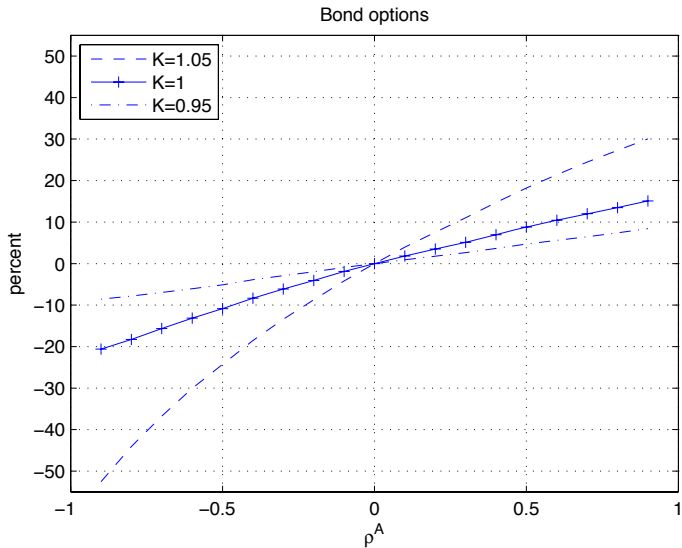
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- As a result, the family of models we have established are very rich in structure, yet are easy to implement.
- In that sense, our analysis complements Duffie-Kan '96.

# Empirical Evidence: Using Kalman Filter



# Importance of Interest Rate-Credit Spread Correlations



# Systemic Credit Shocks and Default Clustering

- The model already allows market-wide events to affect the riskless yield curve and credit spread curves.



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- The model already allows market-wide events to affect the riskless yield curve and credit spread curves.
- We now allow a primary firm's default to impact the credit spread of surviving (secondary) firms.
- Examples: Secondary firm
  - could carry significant debt of the primary firm,
  - may sell much of its goods to a primary firm,
  - may be in competition with the primary firm.

# Credit Spread Dynamics: Secondary Firms

We assume

$$d\lambda_B(t, T) = \mu_B(t, T)dt + \sigma_B(t, T)dz_B(t) + c_{fB}(t, T)dN_f(t) + \sum_{i=1}^{m_B} c_{A_i B}(1 - Y_{A_i}(t))dY_{A_i}(t), \quad \forall t \leq \tau_B,$$

where

- correlation with diffusive riskless term structure:

$$E(dz_f(t)dz'_B(t)) = \Sigma_{m \times n}^B dt = \left( \rho_{ij}^B \right) dt$$

- correlation with firm  $A$ 's diffusive term:

$$E(dz_A(t)dz'_B(t)) = \Sigma_{n \times n}^{AB} dt = \left( \rho_{ij}^{AB} \right) dt$$

- volatility structures are curtailed

## Proposition 3: Pricing Risky Debt of Secondary Firms

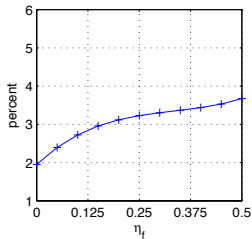
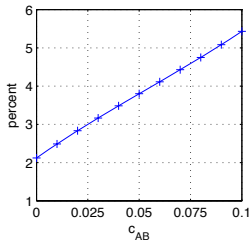
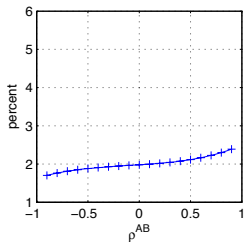
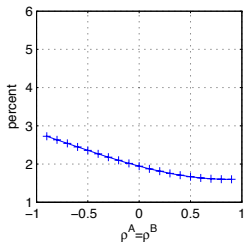
The price of a risky bond issued by a secondary firm is given by  $\Pi_B(t, T) = V_B(t, T)1_{\tau_B > t}$ , where  $V_B(t, T) = P(t, T)S_B(t, T)$  and

$$\begin{aligned}
 S_B(t, T) &= \frac{S_B(0, T)}{S_B(0, t)} e^{-B_0(t, T) - \sum_{j=1}^n (K_{0,j}^B(t, T)\xi_{0,j}^B - K_{1,j}^B(t, T)\xi_{1,j}^B)} \\
 &\times e^{\sum_{i=1}^m \sum_{j=1}^n (K_{2,ij}^B(t, T)\xi_{2,ij}^B - K_{3,ij}^B(t, T)\xi_{3,ij}^B - K_{4,ij}^B(t, T)\xi_{4,ij}^B)} \\
 &\times e^{-K_5^B(t, T)\xi_5^B(t)} \\
 &\times e^{\sum_{i=1}^{m_B} \left( (1 - e^{-c_{A_i B}(T-t)}) U_{A_i B}(t) - c_{A_i B}(T-t) Y_{A_i}(t) \right)}.
 \end{aligned}$$

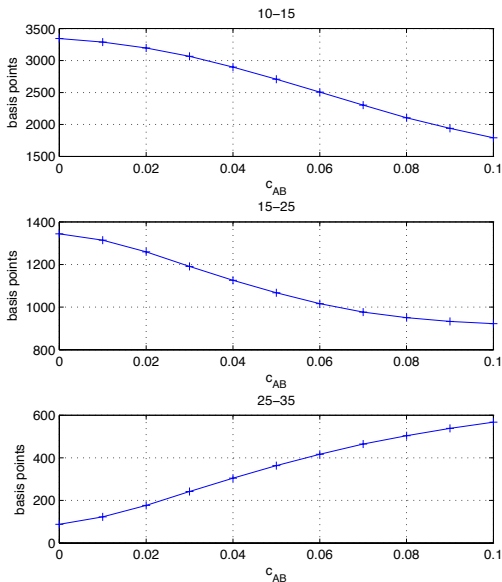
Here,  $B_0(t, T) = \int_t^T \int_0^t g_B(v, u) dv du$ . The  $K^B$  coefficients and  $\xi^B$  state variables are defined as in Proposition 2, and

$$U_{A_i B}(t) = \int_0^{t \wedge \tau_{A_i}} \eta_{A_i}(u) e^{-c_{A_i B}(t-u)} du, \quad \text{for } i = 1, \dots, m_B.$$

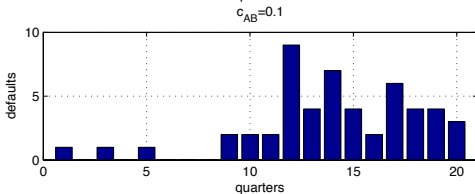
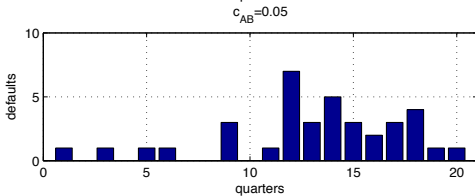
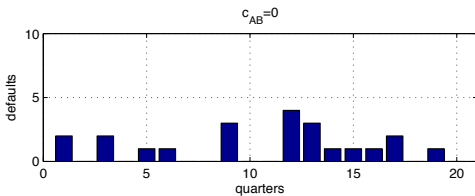
# Importance of Default Contagion: Counterparty Risk in Insurance Contracts



# Importance of Default Contagion: CDS Index Tranches



# Generating Default Clustering



# Importance of the Initial Credit Spread Curve Distr.

Distribution of initial credit spread curves	Tranche spreads		
	10–15	15–25	25–35
$\lambda(0, t) = 0.05$	3451	1450	111
$\lambda(0, 0) \sim \text{Uniform}(0.025, 0.075)$ and $\lambda(0, t) = \lambda(0, 0)$	3523	1528	128
$\lambda(0, 0) \sim \text{Uniform}(0, 0.1)$ and $\lambda(0, t) = \lambda(0, 0)$	3528	1528	121
$\lambda(0, t)$ incr from $\lambda(0, t) = 0.0125$ to $\lambda(0, t) = 0.0875$	2698	1340	107
$\lambda(0, t)$ incr from $\lambda(0, t) = 0.025$ to $\lambda(0, t) = 0.075$	2883	1366	108
$\lambda(0, t)$ incr from $\lambda(0, t) = 0.0375$ to $\lambda(0, t) = 0.0625$	3130	1405	110
$\lambda(0, t) = 0.05$	3451	1450	111
$\lambda(0, t)$ decr from $\lambda(0, t) = 0.0625$ to $\lambda(0, t) = 0.0375$	3857	1514	115
$\lambda(0, t)$ decr from $\lambda(0, t) = 0.075$ to $\lambda(0, t) = 0.025$	4345	1601	120
$\lambda(0, t)$ decr from $\lambda(0, t) = 0.0875$ to $\lambda(0, t) = 0.0125$	4947	1772	141



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- Models have exponentially affine representations for riskless and risky bond prices
  - Yet the variance structures need not be affine
  - The number of state variables is decoupled from the number of stochastic drivers
  - Allow flexible specification of correlations between interest rates and credit spreads
  - Permit default clustering through a variety of channels (diffusive correlations, jumps, contagion effects)