

# Maximum-likelihood regions and smallest credible regions

arXiv:1302.4081[quant-ph]

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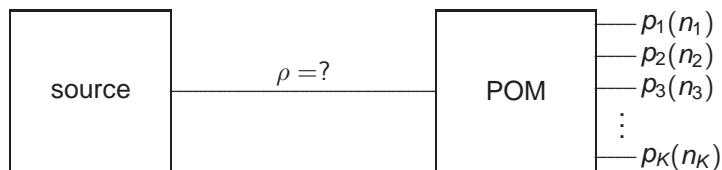
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# Scenario of quantum state estimation



The **source** emits independently and identically prepared quantum-information carriers whose relevant degrees of freedom are described by the “true” statistical operator  $\rho$ , which is unknown.

The **probability-operator measurement** (POM) has  $K$  outcomes  $\Pi_k$  that give rise to the “true” detection probabilities  $p_k$  in accordance with the Born rule,  $p_k = \text{tr}\{\rho\Pi_k\}$ .

The **actual data**  $D$  consist of  $n_1, n_2, \dots, n_K$  detector clicks in *one particular sequence* upon measuring a total of  $N = n_1 + n_2 + \dots + n_K$  copies. [You may want to verify that the sequence is not untypical.]

**State estimation:** Exploit the data for an educated guess about  $\rho = (p_1, p_2, \dots, p_K)$ ; convert  $p \rightarrow \rho$  if you can.

# Principles of quantum state estimation

**1** Be guided by common sense and the methods of classical statistical inference.\*

**2a** Estimate event probabilities from the data, after measuring  $N$  copies.

**2b** Determine the estimator  $\hat{\rho}$  of the state from the estimated probabilities  $\hat{p}_1, \hat{p}_2, \hat{p}_3, \dots$  and, if necessary, invoke additional criteria (such as Jaynes's maximum-entropy criterion).

**Note 1:**  $n = (n_1, n_2, \dots, n_K) \rightarrow \hat{p} = (\hat{p}_1, \dots, \hat{p}_K)$  is what the data tells us;  $\hat{p} \rightarrow \hat{\rho}$  is often not unique, and then the data does *not* tell us  $\hat{\rho}$  and one needs those “additional criteria”.

**Note 2:**  $\hat{p}_k \rightarrow p_k^{(\text{true})}$  for  $N \rightarrow \infty$  (“consistency” — largely a tautology).

\*Read (1) Edwin Jaynes's *Probability Theory — The Logic of Science* and don't ignore his advice; (2) other pertinent statistics literature.

# Reconstruction space (1)

**Reconstruction space**  $\mathcal{R}_0$ : A convex set of  $p$ s such that  $p \leftrightarrow \rho$  is a one-to-one mapping.

**Example 1:** Qubit states  $\rho = \frac{1}{2}(1 + x\sigma_x + y\sigma_y + z\sigma_z)$  measured by the 4-outcome qubit POM with

$$\left. \begin{matrix} p_1 \\ p_2 \end{matrix} \right\} = \frac{1}{4}(1 \pm x), \quad \left. \begin{matrix} p_3 \\ p_4 \end{matrix} \right\} = \frac{1}{4}(1 \pm y)$$

and constraints  $p_1 + p_2 = \frac{1}{2}$ ,  $p_3 + p_4 = \frac{1}{2}$ ,  $p_1^2 + p_2^2 + p_3^2 + p_4^2 \leq \frac{3}{8}$ .

**Example 2:** Qubit states measured by the 3-outcome trine POM with

$$p_1 = \frac{1}{3}(1 + x), \quad \left. \begin{matrix} p_2 \\ p_3 \end{matrix} \right\} = \frac{1}{6}(2 - x \pm \sqrt{3}y)$$

and constraints  $p_1 + p_2 + p_3 = 1$ ,  $p_1^2 + p_2^2 + p_3^2 \leq \frac{1}{2}$

For both examples,  $\mathcal{R}_0$  is the equatorial disk of the Bloch ball; the data provide *no* information about  $z$ .

## Reconstruction space (2)

**Example 3:** Harmonic oscillator measured by the 2-outcome POM with

$$p_1 = \langle 0 | \rho | 0 \rangle, \quad p_2 = 1 - p_1$$

and constraint  $p_1 + p_2 = 1$ .

Here, the reconstruction space consists of all  $\rho = |0\rangle p_1 \langle 0| + p_2 \rho'$  where  $\rho'$  is *any* state with no ground-state component, and the probability space is that of a tossed coin. The data provide only information about the ground-state probability.

### General observations:

- Reconstruction space (may not be unique)  $\equiv$  Probability space
- Because of the quantum constraints, the probability space is usually smaller than that of the  $K$ -sided die:

Quantum State Estimation  
= Classical state estimation with quantum constraints

# Point likelihood, MLE, MLR, SCR

**Point likelihood:**  $L(D|\rho) = p_1^{n_1} p_2^{n_2} \cdots p_K^{n_K}$  = the probability of obtaining data  $D$  if  $\rho$  is the state.

**Maximum-likelihood estimator (MLE)**  $\hat{\rho}_{\text{ML}}$ : That  $\rho$  in  $\mathcal{R}_0$  for which the data are more likely than for any other state:

$$\max_{\rho} L(D|\rho) = L(D|\hat{\rho}_{\text{ML}}).$$

How can we equip the MLE with error bars? Our answer: Use optimal regions.

**Maximum-likelihood region (MLR)**  $\hat{\mathcal{R}}_{\text{ML}}$ : That region of estimators for which the data are more likely than for any other region of the same pre-chosen size.

**Smallest credible region (SCR)**  $\hat{\mathcal{R}}_{\text{SC}}$ : The smallest region with the pre-chosen credibility.

## Size $\equiv$ Prior content

**Scenario 1:** You have a pre-existing notion of size for regions in  $\mathcal{R}_0$ ?  
Fine! Scale all sizes such that  $\mathcal{R}_0$  has unit size; then assign the same prior content to regions of the same size.

**Scenario 2:** You do not have a pre-existing notion of region size?  
Choose the prior of your liking and measure the size of a region by its prior content.

Either way: **Size of a region  $\equiv$  Its prior content.**

**Notation:** The size of region  $\mathcal{R}$  is  $S_{\mathcal{R}} = \int_{\mathcal{R}} (d\rho)$  where  $(d\rho)$  is the prior probability of the infinitesimal space element at state  $\rho$ .

Reference: M.J. Evans, I. Guttman, T. Swartz, *Can. J. Stat.* **34**, 113 (2006).

# MLRs and SCRs are BLRs (1)

1 Joint probability that  $\rho$  is in  $\mathcal{R}$  and data  $D$  is obtained:

$$\text{prob}(D \wedge \mathcal{R}) = \int_{\mathcal{R}} (d\rho) L(D|\rho)$$

2 Prior likelihood  $L(D)$ :  $\text{prob}(D \wedge \mathcal{R}_0) = L(D) = \int_{\mathcal{R}_0} (d\rho) L(D|\rho)$

3 Normalization:  $\sum_D L(D|\rho) = 1$ ,  $\sum_D L(D) = 1$

4 Two factorizations:  $\text{prob}(D \wedge \mathcal{R}) = L(D|\mathcal{R})S_{\mathcal{R}} = C_{\mathcal{R}}(D)L(D)$   
with the **region likelihood**  $L(D|\mathcal{R})$  and the **credibility**  $C_{\mathcal{R}}(D)$ .

Both are conditional probabilities: The region likelihood is the probability of obtaining the data  $D$  if the state is in the region  $\mathcal{R}$ ; the credibility is the probability that the actual state is in the region  $\mathcal{R}$  if the data  $D$  have been obtained—the posterior probability of the region.



## MLRs and SCRs are BLRs (2)

4 Two factorizations:  $\text{prob}(D \wedge \mathcal{R}) = L(D|\mathcal{R})S_{\mathcal{R}} = C_{\mathcal{R}}(D)L(D)$

5 MLR: Maximize the region likelihood for given size,

$$\max_{\mathcal{R}} L(D|\mathcal{R}) = L(D|\hat{\mathcal{R}}_{\text{ML}}) \quad \text{with } S_{\mathcal{R}} = s$$

6 SCR: Minimize the size for given credibility,

$$\min_{\mathcal{R}} S_{\mathcal{R}} = S_{\hat{\mathcal{R}}_{\text{sc}}} \quad \text{with } C_{\mathcal{R}}(D) = c$$

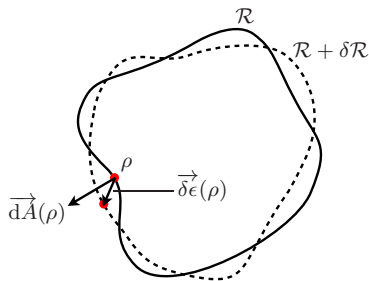
7 These optimization problems are duals of each other:

	MLR	SCR
$S_{\mathcal{R}}$	given	minimize
$\text{prob}(D \wedge \mathcal{R})$	maximize	given

Each MLR is a SCR, each SCR is a MLR.

## MLRs and SCRs are BLRs (3)

8 Infinitesimal variation of region  $\mathcal{R}$  from a distortion of its boundary  $\partial\mathcal{R}$ :



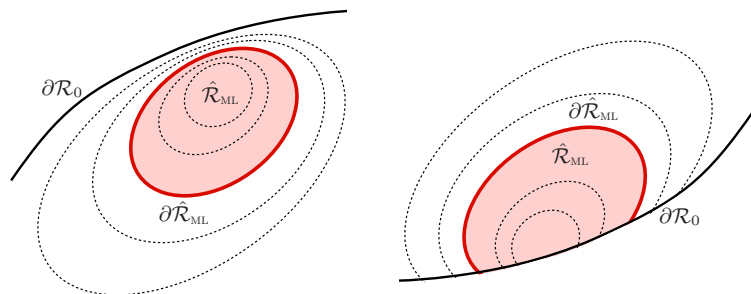
9 Null response of  $S_{\mathcal{R}}$  and  $\text{prob}(D \wedge \mathcal{R})$ :

$$\delta S_{\mathcal{R}} = \int_{\partial\mathcal{R}} \vec{dA}(\rho) \cdot \vec{\delta\epsilon}(\rho) = 0,$$

$$\delta \text{prob}(D \wedge \mathcal{R}) = \int_{\partial\mathcal{R}} \vec{dA}(\rho) \cdot \vec{\delta\epsilon}(\rho) L(D|\rho) = 0$$

## MLRs and SCRs are BLRs (4)

**10** Requiring that both  $\delta S_{\mathcal{R}} = 0$  and  $\delta \text{prob}(D \wedge \mathcal{R}) = 0$  implies that the point likelihood  $L(D|\rho)$  is constant on  $\partial\mathcal{R}$ , and larger inside than on the boundary: The MLRs and the SCRs are bounded-likelihood regions (BLRs), which consist of all  $\rho$ s for which  $L(D|\rho)$  exceeds a threshold value:



Reference: M.J. Evans, I. Guttman, T. Swartz, Can. J. Stat. **34**, 113 (2006).

## MLRs and SCRs are BLRs (5)

**11** The set of BLRs is independent of the prior; each BLR contains the MLE.

**12** Notation:  $\mathcal{R}_\lambda$  is the BLR with  $L(D|\rho) \geq \lambda L(D|\hat{\rho}_{\text{MLE}})$ ;  $s_\lambda = \text{size of } \mathcal{R}_\lambda$ ;  $c_\lambda = \text{credibility of } \mathcal{R}_\lambda$ .

**13** We have

$$c_\lambda > s_\lambda \quad \text{for } 0 < \lambda < 1.$$

In the limit of  $\lambda \rightarrow 1$ , the BLR  $\mathcal{R}_\lambda$  degenerates into the one-point region that contains the MLE, and  $c_\lambda \rightarrow 0$ ,  $s_\lambda \rightarrow 0$ , while

$$\frac{c_\lambda}{s_\lambda} \rightarrow \frac{L(D|\hat{\rho}_{\text{MLE}})}{L(D)} > 1.$$

In the limit of  $\lambda \rightarrow 0$ , the  $\mathcal{R}_\lambda$  becomes full reconstruction space  $\mathcal{R}_0$ , and  $c_\lambda \rightarrow 1$ ,  $s_\lambda \rightarrow 1$ .

# Confidence regions (1)

**MLRs, SCRs:** The data are what they are; the unknown  $\rho$  is regarded as a random variable.

**Confidence regions:** The unknown state is what it is; the data  $D$  (as potentially obtained in many measurements of  $N$  copies each) are regarded as random.

Assign region  $C_D$  to data  $D$ ; the set  $\mathbf{C}$  that is made up of all the  $C_D$ s has the confidence level

$$\gamma(\mathbf{C}) = \min_{\rho} \sum_D \left\{ \begin{array}{ll} L(D|\rho) & \text{if } \rho \text{ is in } C_D \\ 0 & \text{else} \end{array} \right\},$$

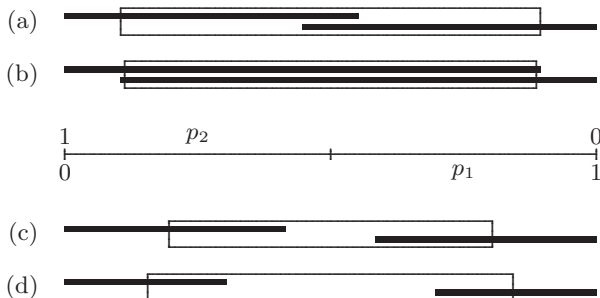
that is: at least the fraction  $\gamma(\mathbf{C})$  of the regions contains  $\rho$  (in many measurements of  $N$  copies each).

**Observations during the 2011 workshop:** Regions with high credibility can be used as confidence regions (Christandl & Renner); a set of BLRs can be a pretty good set of confidence regions (Blume-Kohout).

References: M. Christandl, R. Renner, Phys. Rev. Lett. **109**, 120403 (2012);  
R. Blume-Kohout, arXiv:1202:5270[quant-ph]

## Confidence regions (2)

**Example:** Two copies of the harmonic oscillator measured:



Regions (a) and (b): Two set of confidence regions.

Regions (c): SCRs for the primitive prior  $(d\rho) = dp_1 dp_2 \delta(p_1 + p_2 - 1)$

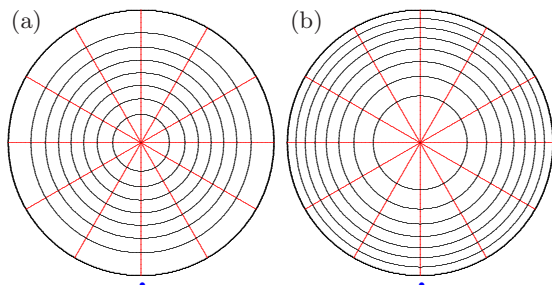
Regions (d): SCRs for Jeffreys prior  $(d\rho) = dp_1 dp_2 \frac{\delta(p_1 + p_2 - 1)}{\pi \sqrt{p_1 p_2}}$

# Choice of prior

- 1 Uniformity — a red herring: All priors are uniform.
- 2 Utility: Be guided by the eventual application.
- 3 Symmetry: Helpful if used with care.
- 4 Invariance — form invariance, really.
- 5 Conjugation: Mock posterior for a target state.
- 6 Marginalization: Convert a prior on the full state space to its marginal on the reconstruction space.

One reference of many: R.E. Kass, L. Wasserman, J. Am. Stat. Assoc. **91**, 1343 (1996)

# Examples of priors, illustrated by uniform tilings (1)

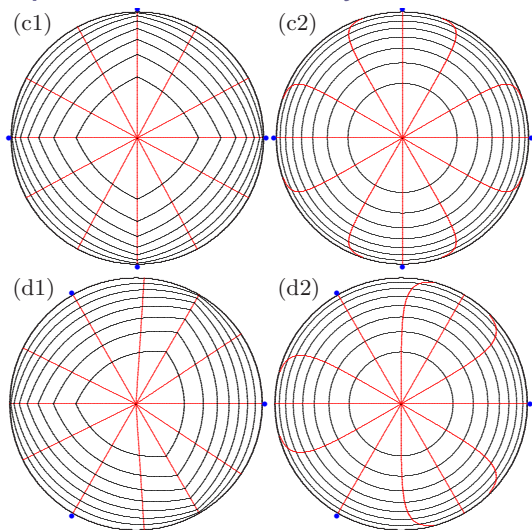


Tiling (a): A prior in the full-qubit space that is rotationally invariant and uniform in the purity, marginalized onto the unit disk.

Tiling (b): The common primitive prior of the 4-outcome POM and the three-outcome POM.



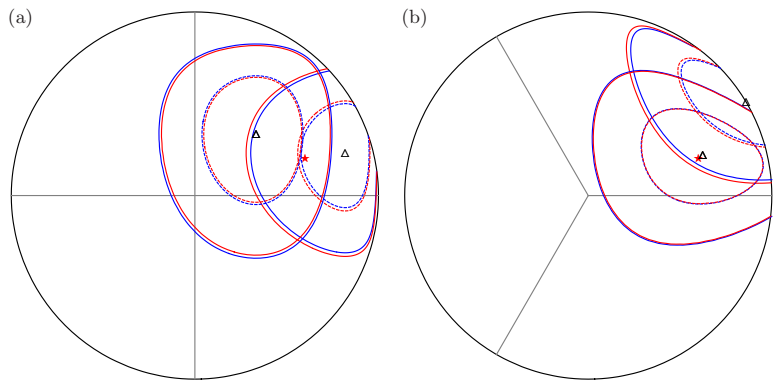
## Examples of priors, illustrated by uniform tilings (2)



Tilings (c1) and (c2): Jeffreys prior for the 4-outcome POM.

Tilings (d1) and (d2): Jeffreys prior for the trine POM.

# Examples of SCRs



SCRs for credibility  $c = 0.5$  and  $c = 0.9$ ; 24 copies measured (in a simulated experiment); primitive (red) and Jeffreys (blue) prior.

(a) 4-outcome POM: counts  $(n_1, n_2, n_3, n_4) = (8, 5, 10, 1)$  and  $(6, 3, 10, 5)$

(b) 3-outcome POM: counts  $(n_1, n_2, n_3) = (15, 8, 1)$  and  $(13, 7, 4)$

# Outlook

**1** While we have efficient methods for calculating the MLE for the data at hand (Many thanks to the Olomouc group!), we are lacking efficient algorithms for finding the SCR.

**2** It may be possible to reduce the dimensionality of the problem if one is really only interested in a few properties of the state (such as the concurrence of a two-qubit state).

**3** For the evaluation of the multi-dimensional integrals, one needs good sampling strategies. Boot strapping of the data may help.

**4** Quantum aspects of the problem enter **only** through the Born rule. Except for the implied restrictions on the probabilities, there is no difference between state estimation in quantum mechanics and statistics. Accordingly, **quantum mechanics can benefit much from methods developed by statisticians.**

Discussions with David Nott (Department of Statistics and Applied Probability, NUS) are gratefully acknowledged.

# THANK YOU