

# Who is Crazier: Bayes or Fisher? A Missing (Data) Perspective on Fiducial Inference

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# Let's Meet the Crazyies

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- Frequentist  $1 - \alpha$  level interval:  $\bar{x} \pm \frac{s}{\sqrt{n}} t_{n-1, 1-\alpha/2}$ .

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**Bayes**

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Fiducial Equation (Taraldsen and Lindqvist 2013)

$$\mathbf{X} = G(\theta, \mathbf{U}) \quad \text{where} \quad \mathbf{X} \in \mathcal{X}, \theta \in \Theta, \mathbf{U} \in \mathcal{U}$$

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**Type of Replication**

**Relevant?**

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- Why should finding  $\mathbf{U}|\mathbf{x}$  be any easier than finding  $\theta|\mathbf{x}$ ?



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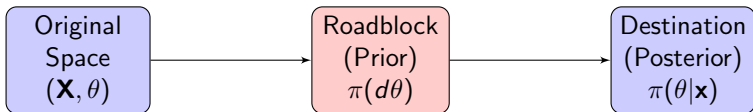
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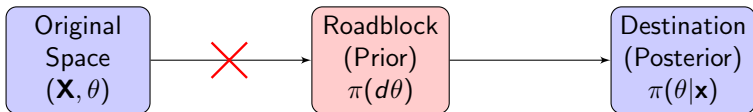
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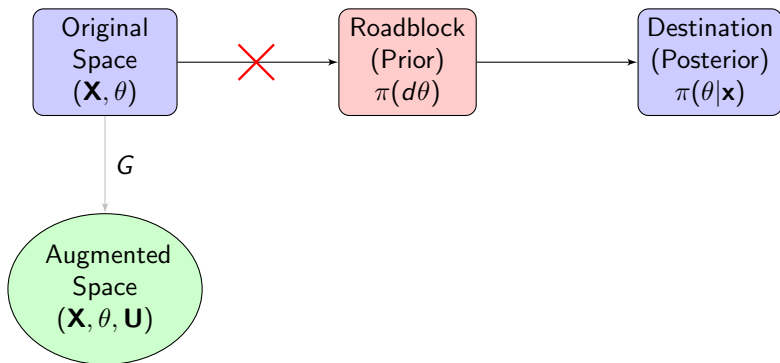
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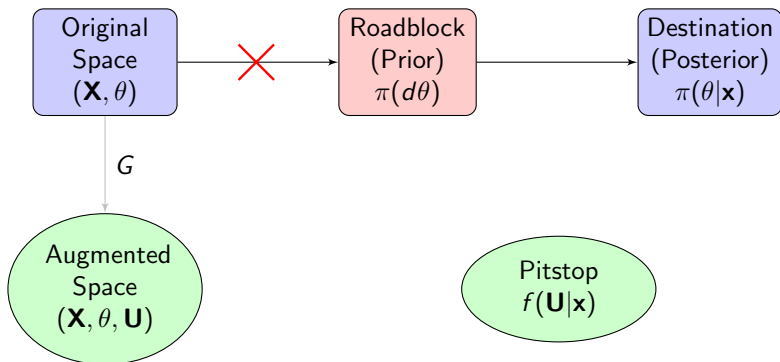
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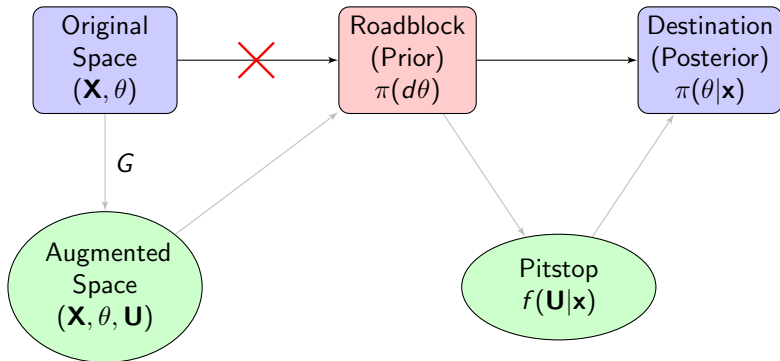
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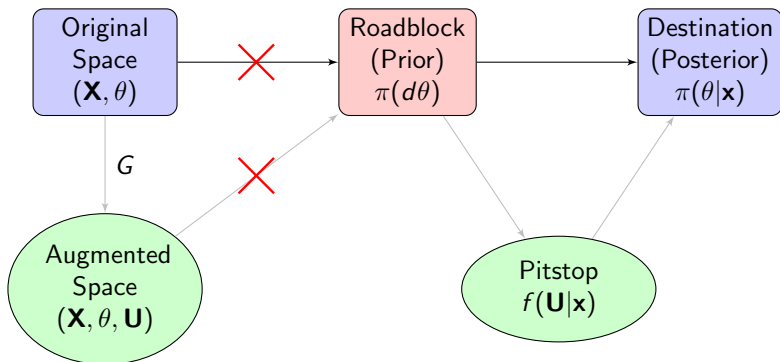
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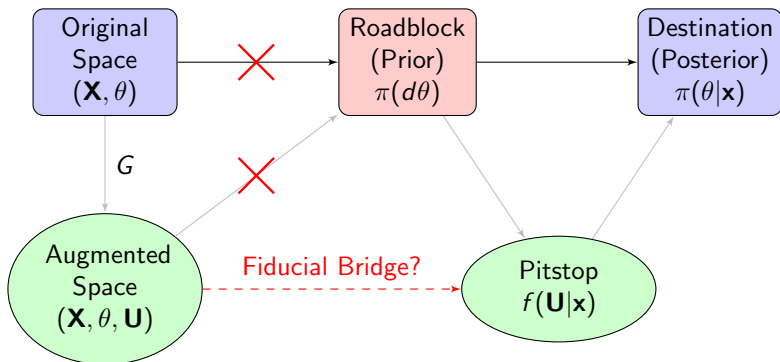
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- **Objective Posterior: Throw away data until we don't need a prior on  $\mu$ .**



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- 1 Obtain  $f(\mathbf{U}|\mathbf{x})$  *without invoking*  $\pi(d\theta)$ .

# An Old Idea

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*The key point is that knowing  $\theta$  is equivalent to knowing  $\mathbf{U}$ ; in other words, inference on  $\theta$  is equivalent to predicting the value of the unobserved  $\mathbf{U}$ . (Martin, Zhang, and Liu 2010)*





# A Missing (Data) Perspective: Prior = Nuisance

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- The conditional distribution of  $\mathbf{X}$  given  $\mathbf{U}$  depends on  $\pi$ .

$$f(\mathbf{x}|\mathbf{U},\pi) = \int f(\mathbf{x}|\mathbf{U},\theta) \pi(d\theta) = \int 1\{\mathbf{x} = G(\theta, \mathbf{U})\} \pi(d\theta)$$

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$$f(\mathbf{U}|\mathbf{x}, \pi) \propto f(\mathbf{U}) f(\mathbf{x}|\mathbf{U}, \pi).$$

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Can we get to  $\pi(\theta|\mathbf{x})$  without going through  $\pi(\theta)$ ?

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Can we predict  $\mathbf{U}$  without any knowledge of the nuisance?



# A Meaningful Definition of Objective

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- $f(\mathbf{x}|\mathbf{U}, \pi)$  contains all information in  $\mathbf{x}$  about  $\mathbf{U}$  and  $\pi$ .

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# Didn't You Say You Were Ancillary?

- Does there exist a statistic,  $A(\mathbf{x})$ , s.t.  $f(A(\mathbf{x})|\mathbf{U})$  is free of the nuisance  $\pi$ ?

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## An Ancillarity Paradox from Basu (1964)

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- Whether or not  $A(\mathbf{x})$  is free of  $\pi$  given  $\mathbf{U}$  depends on  $G$ .



# Representational Ancillarity

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## Definition of $R$ -Ancillarity

A statistic  $A(\mathbf{x})$  is **representationally ancillary w.r.t.  $G$**  if there exists a representation  $A_G$  s.t.  $A(G(\theta, \mathbf{U})) = A_G(\mathbf{U}) \forall \theta$ .

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- $A(\mathbf{x})$  is  $R$ -ancillary if and only if  $\mathbf{U} | A(\mathbf{x}), \pi \sim \mathbf{U} | A(\mathbf{x})$ .

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## Lemma (also see Barnard, 1995)

If  $A_1(\mathbf{x})$  and  $A_2(\mathbf{x})$  are both  $R$ -ancillaries w.r.t.  $G$ , then they are jointly  $R$ -ancillary w.r.t.  $G$ .

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- In Basu's paradox,  $X, Y$  are both ancillaries but they are not  $R$ -ancillaries with respect to the same  $G$ .



# What Do Representations Buy Us?

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# What Do Representations Buy Us?

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- Fiducial inference depends on  $G$  because whether a statistic is ancillary for  $\pi$  depends on whether it is  $R$ -ancillary for  $G$ .

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- We require the notion of *representational* (or functional) *independence*.



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Interested  
In...

regression  
coef.  $\beta$

Fiducial

U

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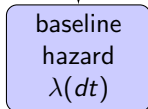
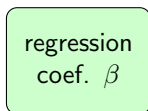
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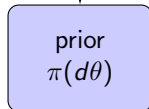
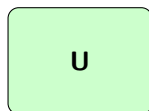
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## Fiducial



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In...

Don't  
Know...

Lossy  
Compress...

## Cox Hazard Model

regression  
coef.  $\beta$

baseline  
hazard  
 $\lambda(dt)$

Full Data  
failure times

## Fiducial

**U**

prior  
 $\pi(d\theta)$

Full Data  
**X**

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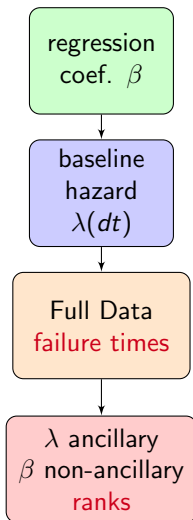
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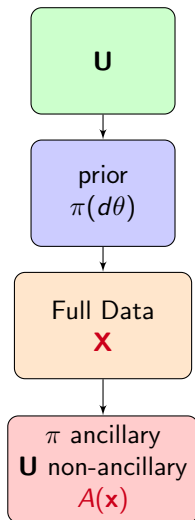
Lossy  
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Inference  
From...

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## Partial Likelihood for $U$

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- $V = \frac{1}{\lambda}E_1$  and  $W = \lambda E_2$  where  $E_1, E_2$  are iid exponential.

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- Why does conditioning on ancillary statistics recover second order information?

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- 
- Why does conditioning on ancillary statistics recover second order information?
  - **Fiducial Answer:** It increases our efficiency of predicting  $\mathbf{U}$ .



# An Often Forgotten Ingredient in Fiducial Cooking

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- $X$  has the following fiducial equation

$$X = G(\theta, U) = \begin{cases} 0 & 0 \leq U < \theta \\ 1 & \theta \leq U < 1 \end{cases}$$

where  $U \sim \text{Unif}[0, 1]$  and  $\theta \in [1/4, 1/2]$ .

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$$X = G(\theta, U) = \begin{cases} 0 & 0 \leq U < \theta \\ 1 & \theta \leq U < 1 \end{cases}$$

where  $U \sim \text{Unif}[0, 1]$  and  $\theta \in [1/4, 1/2]$ .

- **Key Question: Without  $\pi(\theta)$ , what do we know about  $U$ ?  
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# An Often Forgotten Ingredient in Fiducial Cooking

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- **What are we forgetting by doing this?**



# We Observe and Observe that We Observe

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  - ② **How** we came to observe  $\mathbf{U}_0 \in \mathcal{M}(\mathbf{x})$ .
- To correct use the information in (1), **we need to condition on the how**, i.e., condition on the observation process.



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- For  $x \in \mathbb{X}$ , define

$$O_x = \begin{cases} 1 & \text{if } \mathbb{I}^{obs} = \mathbb{I}_x \\ 0 & \text{otherwise} \end{cases}$$



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- $\mathcal{M}(X)$  contains the information  $\mathbb{I}_x$  through  $\mathbb{I}^{obs}$  **and**  $\{O_x\}$ .

$$f(U|\mathcal{M}(X), \pi) = f(U|\mathbb{I}^{obs}, O_0, O_1, \pi)$$

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- Rewrite the posterior using the law of total probability

$$\sum_{t \in \{0,1\}} f(U|\mathbb{I}_x = t, O_0, O_1, \pi) P(\mathbb{I}_x = t|\mathbb{I}^{obs}, O_0, O_1)$$

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- It is this incoherence which leads to the incorrect assumption

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## Confidence Validity (Rubin 1976)

- The posterior  $f(U | U \in \mathcal{M}(x))$  leads to valid confidence regions for  $U$  if the observation process is ignorable

$$P(\mathbb{I}_x = 1 | O_0 = o_0, O_1 = o_1, U, \pi) = P(\mathbb{I}_x = 1 | U)$$



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- Condition for confidence validity does not hold:

$$P(\mathbb{I}_0 = 1 | U, \theta) = \mathbb{I}\{U \leq 1/2\}$$

$$P(\mathbb{I}_0 = 1 | O_0 = 1, O_1 = 0, U, \theta) = \mathbb{I}\{U \leq \theta\}$$

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Conclusions

- Suppose we observe  $\mathbb{I}_0 = 1$  and  $\mathbb{I}_1$  is missing.
- Condition for confidence validity does not hold:

$$P(\mathbb{I}_0 = 1 | U, \theta) = \mathbb{I}\{U \leq 1/2\}$$

$$P(\mathbb{I}_0 = 1 | O_0 = 1, O_1 = 0, U, \theta) = \mathbb{I}\{U \leq \theta\}$$

- **Confidence valid inferences for  $U$  requires modeling  $O_0, O_1$ .**

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- The distribution of  $(O_0, O_1)$  depends on  $\pi$ : the reduction  $X \rightarrow \mathcal{M}(X)$  does not throw away enough information.
- The information  $U \in \mathcal{M}(x)$  seems *free*. But **to use it correctly** requires paying the price of a prior on  $\theta$ .



# The Consequences of Being Cheap

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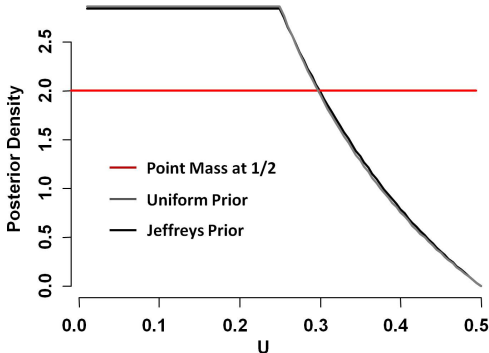
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- **Takeaway:** The information  $\mathbf{U} \in \mathcal{M}(\mathbf{x})$  is *not free*—may require assuming  $\pi$ .

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- **Takeaway:** The information  $\mathbf{U} \in \mathcal{M}(\mathbf{x})$  is *not free*—may require assuming  $\pi$ .
- **Question 1:** So when is this information free? When is  $\mathbf{U}|\mathcal{M}(\mathbf{x})$  objective—free of  $\pi$ ?

- **Takeaway:** The information  $\mathbf{U} \in \mathcal{M}(\mathbf{x})$  is *not free*—may require assuming  $\pi$ .
- **Question 1:** So when is this information free? When is  $\mathbf{U}|\mathcal{M}(\mathbf{x})$  objective—free of  $\pi$ ?
- **Question 2:** What is maximal amount of free information about  $\mathbf{U}$ ? What is the “best” objective posterior for  $\mathbf{U}$ ?



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- $\mathbf{U}|\mathcal{M}(\mathbf{x})$  is an objective posterior for  $\mathbf{U}$  if and only if an  $R$ -ancillary statistic captures the information in  $\mathcal{M}(\mathbf{x})$ .

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## $R$ -Ancillary Regions



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- If  $A(\mathbf{x})$  is  $R$ -ancillary s.t.  $A(G(\theta_0, \mathbf{U}_0)) = A_G(\mathbf{U}_0)$ , the  $R$ -ancillary region defined by  $A$  is the set

$$A_G(\mathbf{U}_0) = \{\mathbf{U} : A_G(\mathbf{U}) = A_G(\mathbf{U}_0)\}$$

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- The smaller the  $R$ -ancillary region, the more informative  $A$  is.
- Any  $R$ -ancillary region is rougher than the  $\mathbf{x}$  compatible region.

$$\mathcal{M}(G(\theta_0, \mathbf{U}_0)) \subset \mathcal{A}_G(\mathbf{U}_0) \quad \forall \theta_0, \forall \mathcal{A}_G$$

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$$\mathcal{M}(G(\theta_0, \mathbf{U}_0)) \subset \mathcal{A}_G(\mathbf{U}_0) \quad \forall \theta_0, \forall \mathcal{A}_G$$

- **What is the smallest we can make  $\mathcal{A}_G$ ? Does a smallest region even exist?**



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Utopia  $\subset$  Cave

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- Ultimately, we hope to restrict  $\mathbf{U}_0$  to Utopia, which is the universal **“Cramer-Rao lower bound for conditioning”**.

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- Can we achieve Utopia?





# Can Statisticians Achieve Utopia?

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## Utopia Can Always Be Achieved

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  - 2 The set of pivot space equivalence classes corresponds to a set of sample space equivalence classes.
  - 3 The index for the sample space equivalence class is observed and is  $R$ -ancillary.
- **Utopia represents an upper bound on the informativeness of an  $R$ -ancillary. It is always achieved.**



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**Classical Way**

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- Subsets of the **sample space**,  $\mathbb{X}$ .

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- Level-sets of ancillary statistics.
- Paradoxes: Which ancillary statistics? Existence?

## New Way

- Subsets of the **pivot space**,  $\mathbb{U}$ .
- Level-sets of  $R$ -ancillary statistics.
- Unique “most relevant” subset—Utopia.

# Making Relevant Subsets Relevant Again

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- Fisher (1934) argued that we should condition on *relevant subsets*. **Of what space?**

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- **Want to give same effort for all difficulties.**



# The Best for U

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# The Best for $\mathbf{U}$

- The **optimal achievable objective** (free of  $\pi$ ) posterior for  $\mathbf{U}$ :  
 $f(\mathbf{U}|A_{\max})$ .

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- The **optimal achievable objective** (free of  $\pi$ ) posterior for  $\mathbf{U}$ :  $f(\mathbf{U}|A_{\max})$ .

“Heaven Is Possible” If and Only If...

With respect to fixed  $G$ ,

$$f(\mathbf{U}|A_{\max}(\mathbf{x})) = f(\mathbf{U}|\mathcal{M}(\mathbf{x})), \quad \forall \mathbf{x} \in \mathbb{X}$$

if and only if  $\forall \mathbf{U} \in \mathbb{U}, \forall \theta \in \Theta, \mathcal{M}(G(\theta, \mathbf{U})) = \mathcal{U}_G(\mathbf{U})$  where  $\mathcal{U}_G(\mathbf{U})$  is the Utopia set containing  $\mathbf{U}$ .

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### Interpretation

- $\mathcal{M}(G(\theta, \mathbf{U}))$  is the information we get back about  $\mathbf{U}$  if data are generated using  $\theta$ .

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### Interpretation

- $\mathcal{M}(G(\theta, \mathbf{U}))$  is the information we get back about  $\mathbf{U}$  if data are generated using  $\theta$ .
- Invariance of  $\mathcal{M}(G(\theta, \mathbf{U}))$  to  $\theta$  implies that information content is **independent** of the true parameter value.



# When Is the Best for $\mathbf{U}$ Good Enough for $\theta$ ?

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## An Objective Posterior for $\theta$

$$\textcircled{1} f(\mathbf{U}|A_{\max}) = f(\mathbf{U}|\mathcal{M}(\mathbf{x})).$$

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## An Objective Posterior for $\theta$

- 1  $f(\mathbf{U}|A_{\max}) = f(\mathbf{U}|\mathcal{M}(\mathbf{x}))$ .
- 2  $\mathbf{x} = G(\theta, \mathbf{U})$  can be solved for  $\theta$ , yielding a function,  $\theta(\mathbf{U}; \mathbf{x})$ , from  $\mathcal{M}(\mathbf{x})$  to  $\Theta$ .

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- $C_\theta$  attains posterior probability  $1 - \alpha$  if (1) and (2) hold.



# Posterior Reflections: From $U$ to $\theta$ Through $G$

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$$\text{Invert } \mathbf{x} = G(\theta, \mathbf{U}) \text{ to} \\ \theta^*(\mathbf{U}; \mathbf{x}) = \{\theta : G(\theta, \mathbf{U}) = \mathbf{x}\}$$

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$$\mathcal{M}(\mathbf{x}) = \mathcal{U}_G(\mathbf{U}_0)$$

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Fiducial posterior  
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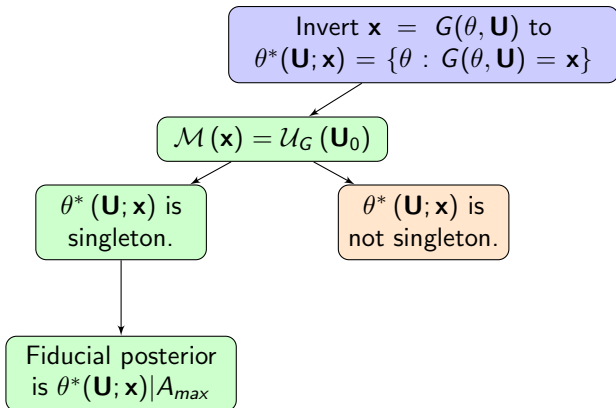
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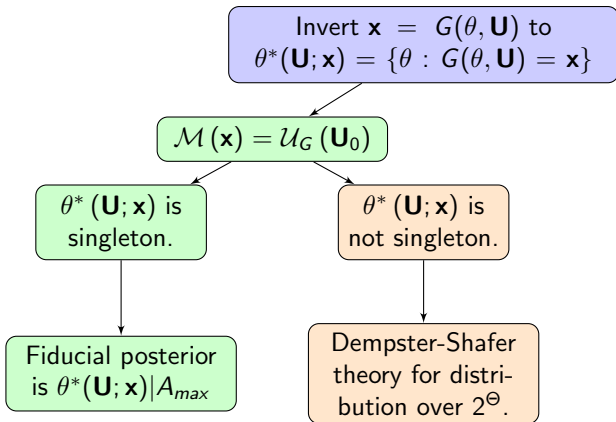
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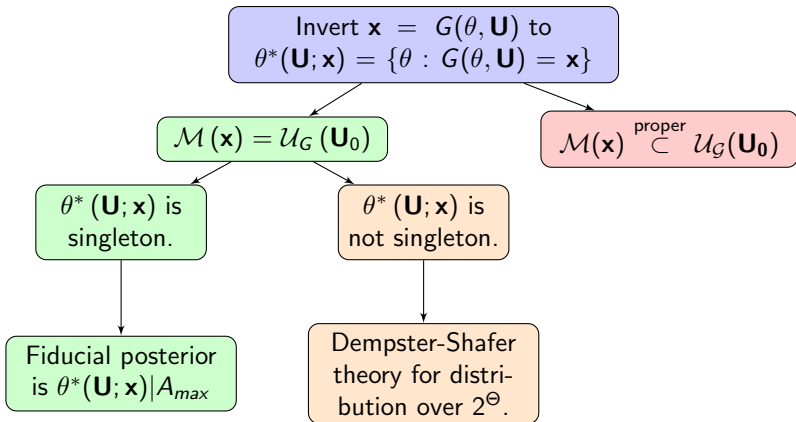
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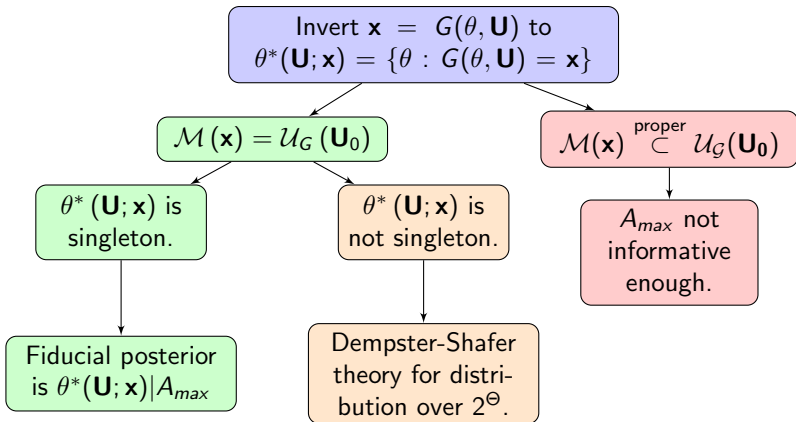
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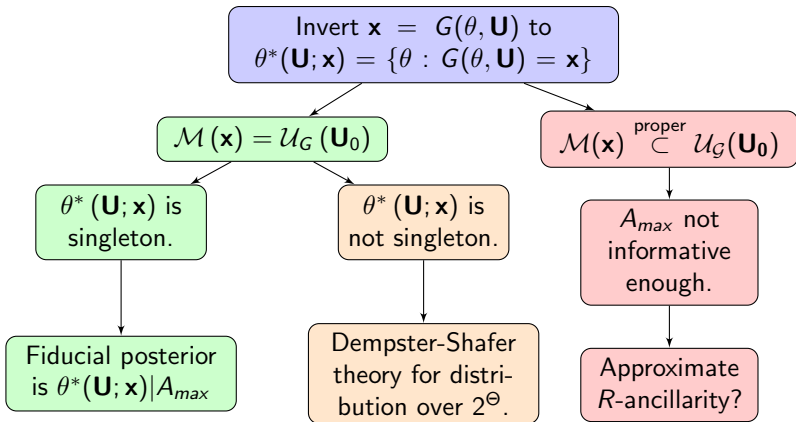
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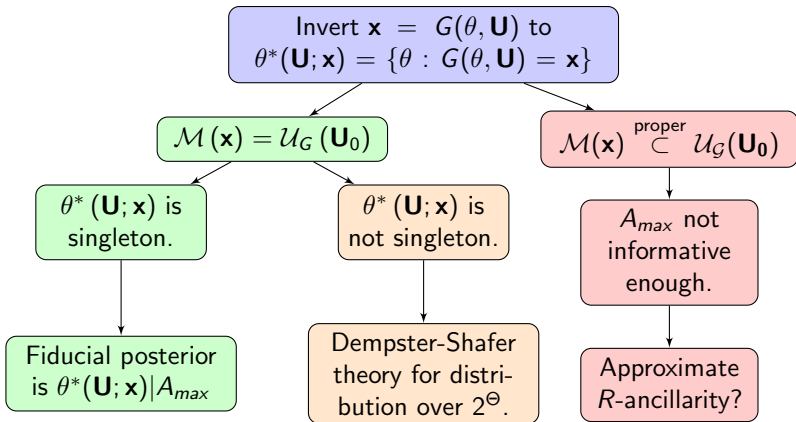
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- **Uncongeniality:** Inconsistent use of the full data,  $\mathbf{x}$ , and the partial data,  $A_{max}$ , in different phases of the analysis.

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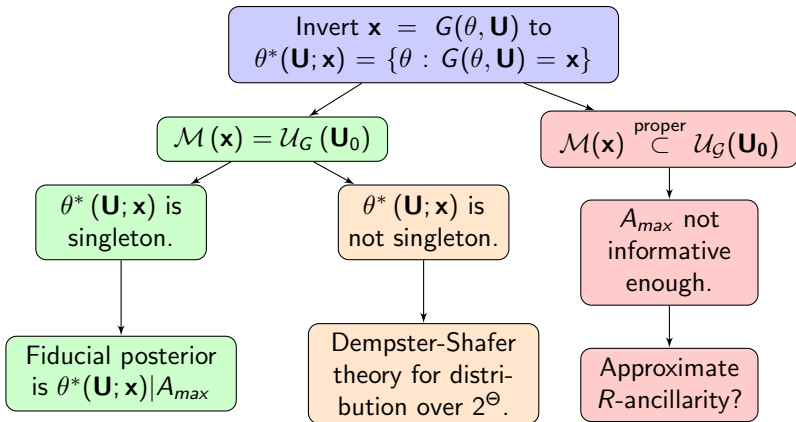
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- **Uncongeniality:** Inconsistent use of the full data,  $\mathbf{x}$ , and the partial data,  $A_{max}$ , in different phases of the analysis.
- **Non-uniqueness:** How does one choose  $G$ ?



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