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# On the free gamma distributions

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## Classical and free infinite divisibility

By  $\mathcal{ID}(*)$  we denote the class of  $*$ -infinitely divisible probability measures on  $\mathbb{R}$ , i.e.

$$
\mu \in \mathcal{ID}(*) \iff \forall n \in \mathbb{N} \; \exists \mu_n \in \mathcal{P}(\mathbb{R}) : \mu = \underbrace{\mu_n * \mu_n * \cdots * \mu_n}_{n \text{ terms}}.
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By  $\mathcal{ID}(\boxplus)$  we denote the class of  $\boxplus$ -infinitely divisible probability measures on  $\mathbb{R}$ , i.e.

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# Classical Lévy-Khintchine representation

**Theorem [Lévy-Khintchine]**. Let  $\mu$  be a probability measure on R and consider its characteristic function

$$
\hat{\mu}(u) = \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i} tu} \, \mu(\mathrm{d} t).
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Then  $\mu$  is infinitely divisible, if and only if  $\hat{\mu}$  has a representation in the form:

$$
\log(\hat{\mu}(u)) = i\eta u - \frac{1}{2}au^2 + \int_{\mathbb{R}} (e^{iut} - 1 - iut1_{[-1,1]}(t)) \; \rho(dt).
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$$

Here  $\eta \in \mathbb{R}$ ,  $a \geq 0$  and  $\rho$  is a Lévy measure on  $\mathbb{R}$ , i.e.

$$
\rho(\{0\})=0, \quad \text{and} \quad \int_{\mathbb{R}} \min\{1,t^2\} \; \rho(\mathrm{d}t) < \infty.
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# The Free Lévy-Khintchine-representation

**Theorem [Bercovici & Voiculescu].** Let  $\mu$  be a probability measure on  $\mathbb R$  with free cumulant transform

$$
\mathcal{C}_{\mu}(z)=zG_{\mu}^{\langle-1\rangle}(z)-1,\qquad (z\in\mathcal{D}(\mu)\subseteq\mathbb{C}^{-}).
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Then  $\mu$  is  $\boxplus$ -infinitely divisible, if and only if  $\mathcal{C}_{\mu}$  has a representation in the form:

$$
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where  $\eta \in \mathbb{R}$ ,  $a \geq 0$  and  $\rho$  is a Lévy measure on  $\mathbb{R}$ .

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where  $\eta \in \mathbb{R}$ ,  $a \geq 0$  and  $\rho$  is a Lévy measure on  $\mathbb{R}$ .

The free characteristic triplet  $(a, \rho, \eta)$  is uniquely determined.

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## The Bercovici-Pata bijection

**Definition.** The Bercovici-Pata bijection  $\Lambda: \mathcal{ID}(*) \to \mathcal{ID}(\boxplus)$  is defined as follows:

$$
\mu \longleftrightarrow \log(\hat{\mu}(u)) = i\eta u - \frac{1}{2}au^2 + \int_{\mathbb{R}} \left( e^{iut} - 1 - iut1_{[-1,1]}(t) \right) \rho(dt)
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\longleftrightarrow (a, \rho, \eta)
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\longleftrightarrow C_{\Lambda(\mu)}(z) = \eta z + az^2 + \int_{\mathbb{R}} \left( \frac{1}{1 - tz} - 1 - tz1_{[-1,1]}(t) \right) \rho(dt)
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Direct formula: For any measure  $\mu$  in  $\mathcal{ID}(*)$  we have

$$
\mathcal{C}_{\Lambda(\mu)}(iz) = \int_0^\infty \log(\hat{\mu}(zx)) e^{-x} dx, \qquad (z < 0).
$$

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## **Properties of the Bercovici-Pata bijection**

(i) If  $\mu_1, \mu_2 \in \mathcal{ID}(*)$ , then  $\Lambda(\mu_1 * \mu_2) = \Lambda(\mu_1) \boxplus \Lambda(\mu_2)$ .

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(ii) If  $\mu \in \mathcal{ID}(*)$  and  $c \in \mathbb{R}$ , then  $\Lambda(D_c\mu) = D_c\Lambda(\mu)$ .

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(iv) For measures  $\mu$ ,  $\mu_1$ ,  $\mu_2$ ,  $\mu_3$ , ... in  $\mathcal{ID}(*)$ , we have

$$
\mu_n\stackrel{w}{\to}\mu\iff \Lambda(\mu_n)\stackrel{w}{\to}\Lambda(\mu).
$$



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## Examples.

(1) Let  $\mu$  be the standard *Gaussian distribution*, i.e.

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\mu(\mathrm{d}t)=\frac{1}{\sqrt{2\pi}}\exp(-\tfrac{1}{2}t^2)\,\mathrm{d}t.
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Then  $\Lambda(\mu)$  is the standard semi-circle distribution, i.e.,

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\mu({n}) = e^{-\lambda} \frac{\lambda^n}{n!}, \qquad (n \in \mathbb{N}_0).
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Then  $\Lambda(\mu)$  is given by

$$
\begin{cases}\n(1-\lambda)\delta_0 + \frac{1}{2\pi t}\sqrt{(t-a)(b-t)} \cdot 1_{[a,b]}(t) \, \mathrm{d}t, & \text{if } 0 \leq \lambda < 1, \\
\frac{1}{2\pi t}\sqrt{(t-a)(b-t)} \cdot 1_{[a,b]}(t) \, \mathrm{d}t, & \text{if } \lambda \geq 1,\n\end{cases}
$$

 $(\overline{\lambda})^2$  and  $b = (1 +$  $(\overline{\lambda})^2$  $(\overline{\lambda})^2$  $(\overline{\lambda})^2$ . where  $a = (1 000$  **[Background](#page-1-0)** [The free exponential distribution](#page-26-0) [Unimodality](#page-59-0) [General free Gamma's](#page-72-0) The free gamma distributions

It is not hard to show that

$$
X \sim \frac{1}{2\pi} \sqrt{4-t^2} 1_{[-1,1]}(t) dt \implies X^2 \sim \frac{1}{4\pi t} \sqrt{t(4-t)} 1_{[0,4]}(t) dt.
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This means that  $X^2$  has the free Poisson distribution with parameter  $\lambda = 1$ .

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#### Natural questions:

- What is  $\Lambda(\chi_1^2)$ ?
- What is Λ(Gamma-distribution)?
- What is  $Λ$ (exponential distribution)?

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## The free exponential distribution

The classical exponential distribution  $\mu(\text{d} x) = \text{e}^{-x}1_{(0,\infty)}(x)\,\text{d} x$  has cumulant function

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Setting  $\nu = \Lambda(\mu)$  we then have for z in  $(-\infty, 0)$  that

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\mathcal{C}_{\nu}(iz) = \int_0^{\infty} \log(\hat{\mu}(zx))e^{-x} dx
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It follows that

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$$
\mathcal{C}_{\nu}(z)=\int_{0}^{\infty}\frac{\mathrm{e}^{-t}}{t}\Big(\frac{1}{1-zt}-1\Big)\,\mathrm{d}t,\ \, \int\limits_{\mathrm{e}^{-\frac{1}{\nu}}\leq \mathrm{e}^{-\frac{1}{\nu}}\leq \mathrm{e}^{-\frac{1}{\nu}}}\mathrm{e}^{-\frac{1}{\nu}}\leq \mathrm{e}^{-\frac{1}{\nu}}\mathrm{e}^{-\frac{1}{\nu}}\leq \mathrm{e}^{-\frac{1}{\nu
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# The free exponential distribution (continued)

Setting  $z=\frac{1}{\omega}$  $\frac{1}{w}$  we find for  $w$  in  $\mathbb{C}^+$  that

$$
\mathcal{C}_{\nu}(1/w) = \int_0^{\infty} \frac{e^{-t}}{t} \left( \frac{1}{1 - t/w} - 1 \right) dt = \int_0^{\infty} \frac{e^{-t}}{t} \left( \frac{w}{w - t} - 1 \right) dt
$$

$$
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so that

$$
G_{\nu}^{\langle -1 \rangle}(\tfrac{1}{w}) = w + wG_{\mu}(w) = w + \int_0^\infty \frac{w}{w - t} e^{-t} dt, \qquad (w \in \mathbb{C}^+).
$$

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## Lebesgue Decomposition

Let  $\mu$  be a (Borel-) probability measure on  $\mathbb R$ , and consider its cumulative distribution function:

$$
F_{\mu}(t)=\mu((-\infty,t]),\qquad(t\in\mathbb{R}),
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as well as its Lebesgue decomposition:

 $\mu = \rho + \sigma$ , where  $\rho \ll \lambda$  and  $\sigma \perp \lambda$ .

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It follows from De la Vallé Poussin's Theorem that

$$
\rho = \mu_{|D_1}, \quad \text{where} \quad D_1 = \left\{ x \in \mathbb{R} \mid \lim_{h \to 0} \frac{F_\mu(x+h) - F_\mu(x)}{h} \text{ exists in } \mathbb{R} \right\}
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\sigma=\mu_{|D_{\infty}}, \quad \text{where} \quad D_{\infty}=\big\{x\in\mathbb{R} \;\big|\; \lim_{h\to 0}\frac{F_{\mu}(x+h)-F_{\mu}(x)}{h}=\infty\big\}.
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\rho = \mu_{|D_1}, \quad \text{where} \quad D_1 = \left\{ x \in \mathbb{R} \mid \lim_{h \to 0} \frac{F_\mu(x+h) - F_\mu(x)}{h} \text{ exists in } \mathbb{R} \right\}
$$

$$
\sigma = \mu_{|D_{\infty}}, \quad \text{where} \quad D_{\infty} = \big\{ x \in \mathbb{R} \; \big| \; \lim_{h \to 0} \frac{F_{\mu}(x+h) - F_{\mu}(x)}{h} = \infty \big\}.
$$

In addition we have that

$$
\rho(\mathrm{d}t)=F_{\mu}'(t)1_{D_1}(t)\,\mathrm{d}t.
$$





It follows then from general theory of Poisson-Stieltjes integrals that

$$
F'_{\mu}(x) = -\frac{1}{\pi} \lim_{y \downarrow 0} \text{Im}(G_{\mu}(x + iy)), \qquad (x \in D_1),
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$$

and that

$$
\lim_{y\downarrow 0} \big|\operatorname{Im}(G_{\mu}(x+{\rm i}y))\big|=\infty, \qquad (x\in D_{\infty}).
$$

In particular the singular part  $\sigma$  of  $\mu$  is concentrated on the set

$$
\big\{x\in\mathbb{R}\,\big|\,\lim_{y\downarrow 0} |G_{\nu_{\alpha}}(x+{\rm i}y)|=\infty\big\}.
$$

 $\leftarrow$ 

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# A fundamental lemma of Bercovici & Voiculescu

For any positive number  $\delta$ , put

$$
\triangle_{\delta} = \{ z \in \mathbb{C}^+ \mid \text{Im}(z) > \delta | \text{Re}(z) | \}.
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If  $\lim_{z\to 0, z\in \Gamma}u(z)=\ell$ , then  $\lim_{z\to 0, z\in \triangle_\delta}u(z)=\ell$  for any positive number  $\delta$ .

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# The free exponential distribution (continued)

We saw before that

$$
G_{\nu}^{\langle -1 \rangle}(\tfrac{1}{w}) = w + wG_{\mu}(w), \qquad (w \in \mathbb{C}^+),
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G_{\nu}^{\langle -1 \rangle}(\tfrac{1}{w}) = w + wG_{\mu}(w), \qquad (w \in \mathbb{C}^+),
$$

so that

$$
\frac{1}{w}=G_\nu(w+wG_\mu(w))
$$

for all  $w$  in  $\mathbb{C}^+$ , such that  $w + wG_\mu(w) \in \mathbb{C}^+.$ 

[Background](#page-1-0) **[The free exponential distribution](#page-26-0)** [Unimodality](#page-59-0) [General free Gamma's](#page-72-0) The curve:  $\int_0^\infty$  $te^{-t}$  $\frac{te^{-t}}{(t-x)^2+y^2} dt = 1.$ 

Let  $c_0$  be the positive constant determined by

$$
\int_0^\infty \frac{t e^{-t}}{(t+c_0)^2} dt = 1, \quad \text{i.e.} \quad c_0 = 0.139688.
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For any x in  $[-c_0, \infty)$  there is a unique positive number  $y = v(x)$ , such that

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\int_0^{\infty} \frac{t e^{-t}}{(t - x)^2 + y^2} dt = 1.
$$

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The curve:  $\int_0^\infty$  $te^{-t}$  $\frac{te^{-t}}{(t-x)^2+y^2} dt = 1.$ 



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# The free exponential distribution (continued)

The free exponential distribution  $\nu$  is absolutely continuous with density given implicitly by

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f_{\nu}(P(x)) = \frac{1}{\pi} \frac{v(x)}{x^2 + v(x)}, \qquad (x \in [-c_0, \infty)),
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where

 $\bullet$ 

$$
P(x) = H(x + i\nu(x)) = \begin{cases} x + 1 + \int_0^\infty \frac{te^{-t}}{x - t} dt, & \text{if } x < -c_0 \\ 2x + 1 - \int_0^\infty \frac{t^2 e^{-t}}{(x - t)^2 + \nu(x)^2} dt, & \text{if } x \ge -c_0 \end{cases}
$$

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- $P(-c_0) \approx 0.1054$ .
- $\nu$  has support [0.1054,  $\infty$ ).

# The density of the free exponential distribution



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# Asymptotic behavior of the free exponential distribution

The asymptotic behavior of  $f_{\nu}(\xi)$  as  $\xi \to \infty$  is given by

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\frac{f_{\nu}(\xi)}{\xi^{-1}e^{\xi}} \longrightarrow e \quad \text{as } \xi \to \infty.
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In particular  $\nu$  has moments of any order.

At the lower bound  $s_0 := \inf \text{supp}(\nu) = P(-c_0)$ , we have that

$$
f_{\nu}(\xi)=\frac{\sqrt{2}}{\pi c_0\sqrt{s_0-c_0^2}}(\xi-s_0)^{1/2}+o(\xi-s_0),\quad\text{as }\xi\downarrow s_0.
$$

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# **Unimodality**

A measure  $\mu$  on  $\mathbb R$  is called *unimodal*, if, for some a in  $\mathbb R$ , it has the form

$$
\mu = \mu({a})\delta_a + f(x) \,\mathrm{d} x,
$$

where f is increasing on  $(-\infty, a)$  and decreasing on  $(a, \infty)$ .

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Question: Are all  $\boxplus$ -selfdecomposable probability measures unimodal?

Theorem [Haagerup+T '11] The free gamma distributions are unimodal.

# Sketch of proof of unimodality

It suffices to show that for any  $\rho$  in  $(0, \infty)$  the equality:

$$
f_{\nu}(\xi)=\rho,\qquad(\xi\in(s_0,\infty))
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has at most 2 solutions,



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\rho = f_{\nu}(P(x)) = -\frac{1}{\pi} \operatorname{Im} \left( \frac{1}{x + \mathrm{i} \nu(x)} \right), \qquad (x \in (-c_0, \infty))
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$$

has at most 2 solutions.

Note that

$$
\left\{z\in\mathbb{C}^+\mid-\frac{1}{\pi}\operatorname{Im}\left(\frac{1}{z}\right)=\rho\right\}=\operatorname{Circle}(\frac{1}{2\pi\rho}i,\frac{1}{2\pi\rho})=:C_{\rho}.
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$$

Hence we want to show that

<span id="page-67-0"></span> $\#(C_{\rho} \cap \text{Graph}(v)) \leq 2.$ 

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## Sketch of proof of unimodality (continued)

In polar coordinates:

$$
\mathcal{C}_{\rho} = \left\{ \frac{1}{\pi \rho} \sin(\theta) e^{i\theta} \mid \theta \in (0, \pi] \right\}.
$$

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# Sketch of proof of unimodality (continued)

In polar coordinates:

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\mathcal{C}_{\rho} = \left\{ \frac{1}{\pi \rho} \sin(\theta) e^{i\theta} \mid \theta \in (0, \pi] \right\}.
$$

Recall also that

Graph(v) = {x + iy 
$$
\in \mathbb{C}^+
$$
 |  $F(x + iy) = 1$ },

where

$$
F(x, y) = \int_0^{\infty} \frac{te^{-t}}{(t - x)^2 + y^2} dt.
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Hence we need to show that

<span id="page-70-0"></span>
$$
\digamma(\tfrac{1}{\pi\rho}\sin(\theta)\mathrm{e}^{\mathrm{i}\theta})=1,\qquad(\theta\in(0,\pi])
$$

has at most 2 solutions.

## Sketch of proof of unimodality (continued)

In polar coordinates:

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<span id="page-71-0"></span>
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$$

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By changes of variables and differentiation under the integral sign, one may verify that the function  $\theta \mapsto F(\frac{1}{\pi n})$  $\frac{1}{\pi\rho}\,\mathsf{sin}(\theta)\mathrm{e}^{\mathrm{i}\theta}\big)$  is strictly decreasing [on](#page-70-0) $(0, \theta_0]$  $(0, \theta_0]$  $(0, \theta_0]$  $(0, \theta_0]$  and st[r](#page-59-0)ictly incr[e](#page-71-0)asing on  $[\theta_0, \pi]$  $[\theta_0, \pi]$  $[\theta_0, \pi]$  [fo](#page-58-0)r [so](#page-72-0)[m](#page-59-0)e  $\theta_0$ [.](#page-84-0)
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### General free Gamma's

The classical Gamma distribution  $\mu_{\alpha}$  with parameter  $\alpha$  is given by

$$
\mu_{\alpha}(\mathrm{d}x) = \frac{1}{\Gamma(\alpha)} x^{\alpha - 1} \mathrm{e}^{-x} 1_{[0,\infty)}(x) \, \mathrm{d}x.
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<span id="page-77-0"></span>**0** 
$$
\lim_{\xi \to \infty} \frac{f_{\alpha}(\xi)}{\xi^{-1} e^{-\xi}} = \alpha e^{\alpha},
$$
 and  $\lim_{\xi \downarrow s_{\alpha}} \frac{f_{\alpha}(\xi)}{\sqrt{\xi - s_{\alpha}}} = \frac{\sqrt{2}}{\pi c_{\alpha} \sqrt{s_{\alpha} - c_{\alpha}^2}}$ .

#### Graphs of  $f_\alpha$  for  $\alpha=\frac{1}{2}$  $\frac{1}{2}$ , 1, 2, 10.



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### Asymptotic behavior as  $\alpha \downarrow 0$

(i) For any  $p$  in  $\mathbb N$  we have that

$$
\frac{1}{\alpha} \int_0^\infty t^p \, \nu_\alpha({\rm d} t) \longrightarrow \int_0^\infty t^{p-1} {\rm e}^{-t} \, {\rm d} t \quad \text{as } \alpha \downarrow 0.
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$$

(ii) For any t in  $(0, \infty)$  we have that

$$
\frac{f_{\alpha}(t)}{\alpha} \longrightarrow t^{-1} e^{-t} \quad \text{as } \alpha \downarrow 0.
$$

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#### Proof of convergence in moments

The (classical) cumulant transform of  $\mu_{\alpha}$  is given by

$$
\log(\hat{\mu}_{\alpha}(u)) = \alpha \int_0^{\infty} (e^{iut} - 1) \frac{e^{-t}}{t} dt = \alpha \sum_{p=1}^{\infty} \frac{i^p (p-1)!}{p!} u^p.
$$

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Hence for any  $p$  in  $\mathbb N$ 

 $r_p(\alpha) := p'$ th free cuml. of  $\nu_\alpha = p'$ th class. cuml. of  $\mu_\alpha = \alpha(p-1)!$ 

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 $r_p(\alpha) := p'$ th free cuml. of  $\nu_\alpha = p'$ th class. cuml. of  $\mu_\alpha = \alpha(p-1)!$ By the moment-cumulant formula it follows that

$$
\int_0^\infty t^p \nu_\alpha(\mathrm{d} t)=r_p(\alpha)+\sum_{k=2}^p\frac{1}{k}\binom{p}{k-1}\sum_{\substack{q_1,\ldots,q_k\geq 1\\q_1+\cdots+q_k=p}}r_{q_1}(\alpha)r_{q_2}(\alpha)\cdots r_{q_k}(\alpha)
$$

= polynomial in  $\alpha$  with no const. term and linear term  $\alpha(p-1)!$ .

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$$

Hence for any  $p$  in  $\mathbb N$ 

 $r_p(\alpha) := p'$ th free cuml. of  $\nu_\alpha = p'$ th class. cuml. of  $\mu_\alpha = \alpha(p-1)!$ By the moment-cumulant formula it follows that

$$
\int_0^\infty t^p \nu_\alpha(\mathrm{d} t)=r_p(\alpha)+\sum_{k=2}^p\frac{1}{k}\binom{p}{k-1}\sum_{\substack{q_1,\ldots,q_k\geq 1\\q_1+\cdots+q_k=p}}r_{q_1}(\alpha)r_{q_2}(\alpha)\cdots r_{q_k}(\alpha)
$$

= polynomial in  $\alpha$  with no const. term and linear term  $\alpha(p-1)!$ . Hence

<span id="page-84-0"></span>
$$
\frac{1}{\alpha}\int_0^\infty t^p\,\nu(\mathrm{d} t)\underset{\alpha\to 0}{\longrightarrow}(p-1)!=\int_0^\infty t^{p-1}\mathrm{e}^{-t}\,\mathrm{d} t.
$$