

Case-base sampling for fitting and validating prognostic models

Workshop on Statistical Issues in Biomarker and Drug Co-development Fields Institute, Toronto

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November 8, 2014

Outline

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• Time matching/risk set sampling (including Cox partial likelihood) eliminates the baseline hazard from the likelihood expression for the hazard ratios.

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- If, however, the absolute risks are of interest, they have to be recovered using the semi-parametric Breslow estimator.
- Alternative approaches for fitting flexible hazard models for estimating absolute risks, not requiring this two-step approach?
- There is; it originates from Mantel (1973) and Hanley & Miettinen (2009).

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- This enables easy fitting of *smooth-in-time* and *non-proportional* hazard models with *multiple time scales*.
- Provides an alternative to Kaplan-Meier-based methods for estimating *discrimination statistics* (e.g. ROC, AUC, risk reclassification probabilities) from *censored survival data*.

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Study base



Case series



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Time matching



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Start again



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Base series



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Age as the time scale





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Base series





Base series matched by the Framingham score





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• The hazard regression can now be fitted using the conditional likelihood expression

$$L(\theta) \equiv \prod_{i=1}^{n} \prod_{t \in (0,\tau]} \left(\frac{\lambda_i(t;\theta)^{\mathrm{d}N_i(t)}}{\rho_i(t) + \lambda_i(t;\theta)} \right)^{\mathrm{d}M_i(t)},$$

where $N_i(t)$ counts the cases, and $M_i(t)$ counts both the case and base series person-moments contributed by individual *i*.

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- This is of logistic regression form with the offset term $\rho_i(t)$ accounting for the base series sampling mechanism.
- Generalizes to multinomial regression when competing causes are present.

Model specification

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Model specification

• Consider the following specification of the hazard function:

$$\begin{split} \lambda_i(t;\theta) &= \exp\{\theta_0 + f_1(t,\theta_1) + f_2(\text{age at baseline}_i + t,\theta_2) \\ &+ f_3(\text{troponin I}_i,\theta_3) \\ &+ \theta_4 \times \text{HDL cholesterol}_i \\ &+ \theta_5 \times \text{non-HDL cholesterol}_i \\ &+ \theta_6 \times \text{treated systolic blood pressure}_i \\ &+ \theta_7 \times \text{untreated systolic blood pressure}_i \\ &+ \theta_8 \times \text{smoker}_i \\ &+ \theta_9 \times \text{prevalent diabetes}_i \}. \end{split}$$

• Here f_1 , f_2 and f_3 are appropriate spline basis functions.

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- The hazard model can be fitted using standard logistic regression procedures.
- The baseline hazard, and consequently, the absolute risk, is obtained as part of the model fit.
- Easy to incorporate multiple time scales and interactions between time and other covariates.
- The time effects themselves can be fitted using flexible specifications, such as regression splines (Hanley & Miettinen, 2009; Saarela & Hanley, 2014).

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- Consequently, we can obtain posterior predictive distributions for discrimination measures such as ROC curves, areas under the curve (AUC), or risk reclassification probabilities.
- Overfitting?
- The procedure works similarly if the risk score has been derived in another sample.

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 Consider for example sensitivity, that is, the probability of the estimated 10-year risk π(X; θ) being at least some threshold risk π*, given the occurrence of the event during the 10 years, and data D:

$$P(\pi(X;\theta) \ge \pi^* \mid N(10) = 1, \theta, D) = \frac{\int_X \mathbf{1}_{\{\pi(x;\theta) \ge \pi^*\}} \pi(x;\theta) P(\mathrm{d}x \mid D)}{\int_X \pi(x;\theta) P(\mathrm{d}x \mid D)}$$

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- The sources of uncertainty here are the unknown parameters θ of the hazard regression model, and the unknown predictive distribution P(X | D) of the prognostic factors.
- If we take $P(dx \mid D) = \sum_{i=1}^{n} \frac{1}{n} \delta_{x_i}(dx)$, a point estimate is given by

$$\frac{\sum_{i=1}^{n} \mathbf{1}_{\{\pi(x_i;\hat{\theta}) \geq \pi^*\}} \pi(x_i;\hat{\theta})}{\sum_{i=1}^{n} \pi(x_i;\hat{\theta})}.$$

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Parametric ROC curves





Kaplan-Meier ROC curves (Heagerty et al. 2000)



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- This corresponds to $P(dx | D) = \sum_{i=1}^{n} w_i \delta_{x_i}(dx)$, where $(w_1, \ldots, w_n) \sim \text{Dirichlet}(1, \ldots, 1)$.

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- This corresponds to $P(dx | D) = \sum_{i=1}^{n} w_i \delta_{x_i}(dx)$, where $(w_1, \ldots, w_n) \sim \text{Dirichlet}(1, \ldots, 1)$.
- The ROC curve and corresponding AUC are calculated at each realization of θ and (w_1, \ldots, w_n) .

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Posterior AUCs for the five models



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- Similarly, this provides an alternative to Kaplan-Meier-based methods for estimating discrimination statistics (e.g. ROC, AUC, risk reclassification probabilities) from censored survival data.
- Bayesian measures of uncertainty can be obtained for these.
- Improving the prediction of CVD in healthy populations, beyond the classic risk factors of CVD, has been challenging.

References

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